

Chapter 6

Independence of random variables

We want to speak of independence of random variables, but all we have is the notion of independence of events. Recall that $X : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}, \mathcal{B})$ is a random variable if $\forall B \in \mathcal{B}$, $X^{-1}(B) \in \mathcal{F}$, or equivalently, from a notational perspective $X^{-1}(\mathcal{B}) \subset \mathcal{F}$. In addition, Example 1.1.4 shows that $X^{-1}(\mathcal{B})$ is a σ -algebra associated with Ω . Hence, it is appropriate to refer to $X^{-1}(\mathcal{B})$ as a sub- σ -algebra of \mathcal{F} . As a matter of terminology, we refer to $X^{-1}(\mathcal{B})$ as the σ -algebra generated by X and it is common to write $\sigma(X) := X^{-1}(\mathcal{B})$.

Before we proceed with the notion of independence of random variables we establish the following theorem, which shows that it is possible to interchange the inverse image and the generation of the σ -algebra of a collection of subsets.

Theorem 6.1. *Let $X : (\Omega, \mathcal{F}, P) \rightarrow (T, \sigma(\mathcal{C}))$ and \mathcal{C} a class of subsets of T . Then,*

$$X^{-1}(\sigma(\mathcal{C})) = \sigma(X^{-1}(\mathcal{C})).$$

Proof. From Example 1.1.4 $X^{-1}(\sigma(\mathcal{C}))$ is a σ -algebra associated with Ω . Since $\mathcal{C} \subset \sigma(\mathcal{C})$, $X^{-1}(\mathcal{C}) \subset X^{-1}(\sigma(\mathcal{C}))$ and consequently $\sigma(X^{-1}(\mathcal{C})) \subset X^{-1}(\sigma(\mathcal{C}))$.

For the reverse of the last set containment, first define $\mathcal{U} := \{U \in 2^T : X^{-1}(U) \in \sigma(X^{-1}(\mathcal{C}))\}$. By definition of \mathcal{U} , $X^{-1}(U) \subset \sigma(X^{-1}(\mathcal{C}))$ and if $C \in \mathcal{C}$, $X^{-1}(C) \in \sigma(X^{-1}(\mathcal{C}))$. Hence, $\mathcal{C} \subset \mathcal{U}$ and $X^{-1}(\mathcal{C}) \subset X^{-1}(\mathcal{U}) \subset \sigma(X^{-1}(\mathcal{C}))$. Then, if \mathcal{U} is a σ -algebra we have that $\sigma(\mathcal{C}) \subset \mathcal{U}$ and $X^{-1}(\sigma(\mathcal{C})) \subset X^{-1}(\mathcal{U}) \subset \sigma(X^{-1}(\mathcal{C}))$.

We now show that \mathcal{U} is a σ -algebra. First, $T \in \mathcal{U}$ since $X^{-1}(T) = \Omega \in \sigma(X^{-1}(\mathcal{C}))$. Second, if $U \in \mathcal{U}$, then $X^{-1}(U) \in \sigma(X^{-1}(\mathcal{C}))$ and $(X^{-1}(U))^c = X^{-1}(U^c) \in \sigma(X^{-1}(\mathcal{C}))$, showing that $U^c \in \mathcal{U}$. Third, if $\{U_i\}_{i \in \mathbb{N}} \subset \mathcal{U}$, then $X^{-1}(U_i) \in \sigma(X^{-1}(\mathcal{C}))$ for all $i \in \mathbb{N}$, which implies that $\bigcup_{i \in \mathbb{N}} X^{-1}(U_i) = X^{-1}\left(\bigcup_{i \in \mathbb{N}} U_i\right) \in \sigma(X^{-1}(\mathcal{C}))$, showing that $\bigcup_{i \in \mathbb{N}} U_i \in \mathcal{U}$, and completing the proof. ■

Earlier, we defined a finite collection of events $\{E_i\}_{i=1}^n \subset \mathcal{F}$ for $2 \leq n \in \mathbb{N}$ as being independent if

$$P\left(\bigcap_{j \in J} E_j\right) = \prod_{j \in J} P(E_j) \text{ for any } J \subset I = \{1, \dots, n\}. \quad (6.1)$$

We extend this definition of independence to sub- σ -algebras of a probability space and to random elements.

Definition 6.1. *Let $2 \leq n \in \mathbb{N}$ and $\{\mathcal{C}_i\}_{i=1}^n$ be a collection of classes of events. That is, each \mathcal{C}_i contains events associated with the probability space (Ω, \mathcal{F}, P) . The collection $\{\mathcal{C}_i\}_{i=1}^n$ is said to be independent if for any $E_i \in \mathcal{C}_i$ we have that $\{E_i\}_{i=1}^n$ is an independent collection of events.*

This definition motivates the following:

Definition 6.2. *Let $I = \{1, \dots, n\}$, $2 \leq n \in \mathbb{N}$ and (Ω, \mathcal{F}, P) be a probability space. Then,*

(a) *Sub- σ -algebras \mathcal{F}_i of \mathcal{F} with $i \in I$ are independent if for every $J \subset I$ and all $E_i \in \mathcal{F}_i$*

$$P\left(\bigcap_{j \in J} E_j\right) = \prod_{j \in J} P(E_j),$$

(b) *Random variables $X_i : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}, \mathcal{B})$ for $i \in I$ are independent if the sub σ -algebras $\sigma(X_i) := X_i^{-1}(\mathcal{B})$ are independent.*

As a matter of notation, whenever two σ -algebras \mathcal{F}_1 and \mathcal{F}_2 are independent we write $\mathcal{F}_1 \perp\!\!\!\perp \mathcal{F}_2$. Similarly, whenever two random variables X_1 and X_2 are independent we write $X_1 \perp\!\!\!\perp X_2$.

Remark 6.1. Recall that by definition $X^{-1}(B) = \{\omega : X(\omega) \in B\}$. Hence, when we write $P(X \in B)$ we mean $P(X^{-1}(B))$, for $B \in \mathcal{B}$.

The following theorem provides a criterion for establishing the independence of σ -algebras.

Theorem 6.2. Let (Ω, \mathcal{F}, P) be a probability space. For $i = 1, \dots, n$ and $n \in \mathbb{N}$ let \mathcal{C}_i be a non-empty collection of events satisfying:

1. \mathcal{C}_i is a π -system,
2. $\{\mathcal{C}_i\}_{i=1}^n$ is an independent collection.

Then, $\{\sigma(\mathcal{C}_i)\}_{i=1}^n$ is an independent collection.

Proof. First, let $n = 2$. In this case we need to consider \mathcal{C}_1 and \mathcal{C}_2 . Choose an arbitrary $A_2 \in \mathcal{C}_2$ and let $\mathcal{E} = \{A \in \mathcal{F} : P(A \cap A_2) = P(A)P(A_2)\}$. \mathcal{E} is the collection of events that are independent of A_2 . Now, note that:

1. $P(\Omega \cap A_2) = P(A_2) = P(\Omega)P(A_2)$ since $P(\Omega) = 1$. Thus, $\Omega \in \mathcal{E}$.
2. Suppose $A \in \mathcal{E}$. Note that

$$\begin{aligned} P(A^c \cap A_2) &= P((\Omega - A) \cap A_2) = P(A_2 - (A \cap A_2)) = P(A_2) - P(A \cap A_2) \\ &= P(A_2) - P(A)P(A_2) \text{ since } A \in \mathcal{E} \\ &= P(A_2)(1 - P(A)) = P(A_2)P(A^c). \end{aligned}$$

Thus, if $A \in \mathcal{E}$ we have that $A^c \in \mathcal{E}$.

3. If $\{A_n\}_{n \in \mathbb{N}} \subset \mathcal{E}$ is a pairwise disjoint collection

$$\begin{aligned} P\left(\left(\bigcup_{n \in \mathbb{N}} A_n\right) \cap A_2\right) &= P\left(\bigcup_{n \in \mathbb{N}} (A_n \cap A_2)\right) \\ &= \sum_{n \in \mathbb{N}} P(A_n \cap A_2) \text{ since the sets in the union are disjoint} \\ &= \sum_{n \in \mathbb{N}} P(A_n)P(A_2) = P(A_2)P\left(\bigcup_{n \in \mathbb{N}} A_n\right) \text{ since the sets } A_n \text{ are in } \mathcal{E}. \end{aligned}$$

Thus, if $\{A_n\}_{n \in \mathbb{N}} \subset \mathcal{E}$ is a pairwise disjoint collection, we have that $\bigcup_{n \in \mathbb{N}} A_n \in \mathcal{E}$.

Since 1-3 are the defining properties of a Dynkin system, we conclude that \mathcal{E} is a Dynkin system. Note also that, by assumption, \mathcal{C}_1 is independent of \mathcal{C}_2 , every $A_1 \in \mathcal{C}_1$ is in \mathcal{E} . Thus, $\mathcal{C}_1 \subset \mathcal{E}$. By Theorem [2.3](#), since \mathcal{C}_1 is a π -system $\mathcal{E} \supset \delta(\mathcal{C}_1) = \sigma(\mathcal{C}_1)$.¹ Thus, all the events in $\sigma(\mathcal{C}_1)$ are in \mathcal{E} and we can conclude that $\sigma(\mathcal{C}_1)$ is independent of \mathcal{C}_2 . We can also conclude, by the symmetry of the argument, that $\sigma(\mathcal{C}_2)$ is independent of \mathcal{C}_1 .

Now, repeat the argument above by choosing an arbitrary $A_2 \in \sigma(\mathcal{C}_2)$. Then, \mathcal{E} is a Dynkin system, and by the fact that $\sigma(\mathcal{C}_2)$ is independent of \mathcal{C}_1 we have that $\mathcal{C}_1 \subset \mathcal{E}$ and, as above, $\sigma(\mathcal{C}_1) \subset \mathcal{E}$. Consequently, $\sigma(\mathcal{C}_1)$ is independent of $\sigma(\mathcal{C}_2)$.

If $n = 3$, we need to consider $\sigma(\mathcal{C}_1)$, $\sigma(\mathcal{C}_2)$, $\sigma(\mathcal{C}_3)$ and establish that for any $A_i \in \sigma(\mathcal{C}_i)$:
1. $P(\bigcap_{i=1}^3 A_i) = \prod_{i=1}^3 P(A_i)$, 2. $P(A_i \cap A_j) = P(A_i)P(A_j)$ for $1 \leq i < j \leq 3$. By assumption $\{\mathcal{C}_i\}_{i=1}^3$ forms an independent collection. Hence, using the arguments for $n = 2$, the conditions in item 2 are met. We now verify item 1.

Fix $A_2 \in \mathcal{C}_2$, $A_3 \in \mathcal{C}_3$ and consider $\mathcal{E} = \{A_1 \in \mathcal{F} : P(\bigcap_{i=1}^3 A_i) = \prod_{i=1}^3 P(A_i)\}$. Then,

$$\begin{aligned} P(\Omega \cap A_2 \cap A_3) &= P(A_2 \cap A_3) \\ &= P(A_2)P(A_3) \text{ by independence of } \mathcal{C}_2 \text{ and } \mathcal{C}_3 \\ &= P(\Omega)P(A_2)P(A_3) \text{ since } P(\Omega) = 1. \end{aligned}$$

Hence, we conclude that $\Omega \in \mathcal{E}$. Now, let $A \in \mathcal{E}$. Then,

$$\begin{aligned} P(A^c \cap A_2 \cap A_3) &= P(A_2 \cap A_3) - P(A \cap A_2 \cap A_3) \\ &= P(A_2 \cap A_3) - P(A)P(A_2)P(A_3) \text{ since } A \in \mathcal{E} \\ &= P(A_2)P(A_3) - P(A)P(A_2)P(A_3) \text{ by independence of } \mathcal{C}_2 \text{ and } \mathcal{C}_3 \\ &= P(A^c)P(A_2)P(A_3). \end{aligned}$$

¹ $\delta(\mathcal{C}_1)$ is the smallest Dynkin system generated by \mathcal{C}_1 .

Hence, $A^c \in \mathcal{E}$. Now, let $\{A_n\}_{n \in \mathbb{N}} \subset \mathcal{E}$ be pairwise disjoint. Then,

$$\begin{aligned} P \left(\left(\bigcup_{n \in \mathbb{N}} A_n \right) \cap A_2 \cap A_3 \right) &= P \left(\bigcup_{n \in \mathbb{N}} (A_n \cap A_2 \cap A_3) \right) \\ &= \sum_{n \in \mathbb{N}} P(A_n \cap A_2 \cap A_3) \text{ since } \{A_n\}_{n \in \mathbb{N}} \text{ is pairwise disjoint} \\ &= \sum_{n \in \mathbb{N}} P(A_n)P(A_2)P(A_3) \text{ since } A_n \in \mathcal{E}. \end{aligned}$$

Hence, $\bigcup_{n \in \mathbb{N}} A_n \in \mathcal{E}$. Thus, \mathcal{E} is a Dynkin system. Since $\{\mathcal{C}_i\}_{i=1}^3$ is an independent collection, $A_1 \in \mathcal{C}_1 \implies A_1 \in \mathcal{E}$. Hence, $\sigma(\mathcal{C}_1) = \delta(\mathcal{C}_1) \subset \mathcal{E}$. Thus, $\{\sigma(\mathcal{C}_1), \mathcal{C}_2, \mathcal{C}_3\}$ forms an independent collection. Now, fix $A_1 \in \sigma(\mathcal{C}_1)$ and $A_2 \in \mathcal{C}_2$. Define $\mathcal{E} = \{A_3 \in \mathcal{F} : P(A_3 \cap A_1 \cap A_2) = P(A_3)P(A_1)P(A_2)\}$. \mathcal{E} is a Dynkin system and $A_3 \in \mathcal{C}_3 \implies A_3 \in \mathcal{E}$ since $\{\sigma(\mathcal{C}_1), \mathcal{C}_2, \mathcal{C}_3\}$ forms an independent collection. Hence, $\sigma(\mathcal{C}_3) = \delta(\mathcal{C}_3) \subset \mathcal{E}$. Thus, $\{\sigma(\mathcal{C}_1), \mathcal{C}_2, \sigma(\mathcal{C}_3)\}$ forms an independent collection. Lastly, fix $A_1 \in \sigma(\mathcal{C}_1)$ and $A_3 \in \sigma(\mathcal{C}_3)$. Define $\mathcal{E} = \{A_2 \in \mathcal{F} : P(A_2 \cap A_1 \cap A_3) = P(A_2)P(A_1)P(A_3)\}$. \mathcal{E} is a Dynkin system and $A_2 \in \mathcal{C}_2 \implies A_2 \in \mathcal{E}$ since $\{\sigma(\mathcal{C}_1), \mathcal{C}_2, \sigma(\mathcal{C}_3)\}$ forms an independent collection. Hence, $\sigma(\mathcal{C}_2) = \delta(\mathcal{C}_2) \subset \mathcal{E}$. Thus, $\{\sigma(\mathcal{C}_i)\}_{i=1}^3$ forms an independent. Repeated use of this argument establishes that $\{\sigma(\mathcal{C}_i)\}_{i=1}^n$ is an independent collection. ■

Definition 6.3. Let \mathcal{I} be an arbitrary index set (not necessarily finite or even countable). The collection $\{\mathcal{C}_i\}_{i \in \mathcal{I}}$ is independent if for each finite $I \subset \mathcal{I}$, the collection $\{\mathcal{C}_i\}_{i \in I}$ is independent.

Given Definition [6.3](#), we have the following corollary to Theorem [6.2](#).

Corollary 6.1. Let $\{\mathcal{C}_i\}_{i \in \mathcal{I}}$ be a collection of non-empty independent π -systems. Then, $\{\sigma(\mathcal{C}_i)\}_{i \in \mathcal{I}}$ is an independent collection.

Theorem 6.3. Let $\{\mathcal{F}_i\}_{i \in \mathcal{I}}$ be an independent collection of σ -algebras, S be an index set, $\{I_s\}_{s \in S}$ be a pairwise disjoint collection of subsets $I_s \subset \mathcal{I}$ and

$$\mathcal{F}_{I_s} = \sigma \left(\bigcup_{i \in I_s} \mathcal{F}_i \right).$$

Then, $\{\mathcal{F}_{I_s}\}_{s \in S}$ is an independent collection of σ -algebras.

Proof. Suppose S is finite. Define $C_{I_s} := \left\{ \bigcap_{t \in T} B_t : B_t \in \mathcal{F}_t, T \subset I_s, T \text{ finite} \right\}$. If $C, C' \in C_{I_s}$, then

$$C \cap C' = \left(\bigcap_{t \in T \subset I_s} B_t \right) \cap \left(\bigcap_{t \in T' \subset I_s} B'_t \right) = \bigcap_{t \in T \cup T'} (B_t \cap B'_t) \text{ where } B_t, B'_t \in \mathcal{F}_t.$$

But since \mathcal{F}_t is a σ -algebra, $B_t \cap B'_t \in \mathcal{F}_t$, establishing that C_{I_s} is a π -system. Now, let $C \in C_{I_s}$ and $C' \in C_{I_{s'}}$ for any $s \neq s', s, s' \in S$. Then,

$$P(C \cap C') = P \left(\left(\bigcap_{t \in T \subset I_s} B_t \right) \cap \left(\bigcap_{t \in T' \subset I_{s'}} B'_t \right) \right) = \prod_{t \in T \subset I_s} P(B_t) \prod_{t \in T' \subset I_{s'}} P(B'_t) = P(C)P(C'),$$

where the penultimate equality follows from independence of $\{\mathcal{F}_i\}_{i \in \mathcal{I}}$ and the fact that $\{I_s\}_{s \in S}$ is a pairwise disjoint collection. The last equality follows from independence of $\{\mathcal{F}_i\}_{i \in \mathcal{I}}$ and the definition of C and C' . Hence, $\{C_{I_s}\}_{s \in S}$ forms an independent collection of π -systems.

Now, note that if $C \in C_{I_s}$, $C = \bigcap_{t \in T} B_t$ where $B_t \in \mathcal{F}_t, T \subset I_s$. Hence, $B_t \in \bigcup_{t \in T} \mathcal{F}_t \subset \bigcup_{t \in I_s} \mathcal{F}_t$ and consequently $B_t \in \mathcal{F}_{I_s} = \sigma \left(\bigcup_{s \in I_s} \mathcal{F}_s \right)$. But since \mathcal{F}_{I_s} is a σ -algebra, it is closed under countable intersections, and we conclude that $C \in \mathcal{F}_{I_s}$. Hence, $C_{I_s} \subset \mathcal{F}_{I_s}$ and

$$\sigma(C_{I_s}) \subset \mathcal{F}_{I_s} \tag{6.2}$$

Also, setting $T = \{t\}$, we have that C_{I_s} contains B_t for $t \in I_s$. Hence, $\sigma(C_{I_s}) \supset B_t$ for $t \in I_s$.

Since, σ -algebras are closed under countable unions, $\sigma(C_{I_s}) \supset \bigcup_{t \in I_s} B_t$ and we conclude that

$$\sigma(C_{I_s}) \supset \sigma \left(\bigcup_{t \in I_s} B_t \right) = \mathcal{F}_{I_s}. \tag{6.3}$$

By the set containments in [6.2](#) and [6.3](#) we conclude that $\sigma(C_{I_s}) = \mathcal{F}_{I_s}$. Hence, by Theorem [6.2](#) we have that $\{\mathcal{F}_{I_s}\}_{s \in S}$ is an independent collection of σ -algebras. Given Definition [6.3](#) and Corollary [6.1](#) the index set S can be arbitrary, completing the proof. \blacksquare

Example 6.1. Consider a collection $\{X_i\}_{i \in \mathbb{N}}$ of independent random variables. Then, by definition $\{\sigma(X_i)\}_{i \in \mathbb{N}}$ is an independent collection of σ -algebras. Let $S = \{1, 2\}$, $I_1 = \{1, \dots, n\}$ and $I_2 = \{n+1, n+2, \dots\}$. Then,

$$\sigma(X_1, \dots, X_n) := \sigma\left(\bigcup_{i=1}^n \mathcal{F}_i\right) \perp\!\!\!\perp \sigma\left(\bigcup_{i>n} \mathcal{F}_i\right) := \sigma(X_{n+1}, \dots).$$

Since for all $i = 1, \dots, n$, X_i is such that $\sigma(X_i) \subset \sigma(X_1, \dots, X_n)$, these random variables are all $\sigma(X_1, \dots, X_n) - \mathcal{B}$ measurable. Hence, $\sum_{i=1}^n X_i$ is $\sigma(X_1, \dots, X_n) - \mathcal{B}$ measurable. Similarly, for any $n_1 \in \mathbb{N}$, $\sum_{i=n+1}^{n+n_1} X_i$ is $\sigma(X_{n+1}, \dots, X_{n+n_1}) - \mathcal{B}$ measurable, and consequently

$$\sum_{i=1}^n X_i \perp\!\!\!\perp \sum_{i=n+1}^{n+n_1} X_i.$$

Definition [6.2](#) can be naturally expanded in accordance to Definition [6.3](#) to accommodate an arbitrarily indexed collection of random variables. We now provide some characterizations for independence of random variables.

Definition 6.4. Let $\{X_i\}_{i \in \mathcal{I}}$ be a collection of random variables defined on the probability space (Ω, \mathcal{F}, P) . For any finite $I \subset \mathcal{I}$, the finite dimensional distribution function (fddf) is given by

$$F_I(x_i, i \in I) = P\left(\bigcap_{i \in I} \{\omega : X_i(\omega) \leq x_i\}\right) \text{ for } x_i \in \mathbb{R}. \quad (6.4)$$

Theorem 6.4. The collection $\{X_i\}_{i \in \mathcal{I}}$ of random variables defined on the probability space (Ω, \mathcal{F}, P) is independent if, and only if, for all finite $I \subset \mathcal{I}$ we have

$$F_I(x_i, i \in I) = \prod_{i \in I} P(\{\omega : X_i(\omega) \leq x_i\}) \text{ for } x_i \in \mathbb{R}. \quad (6.5)$$

Proof. From Definition [6.3](#) it suffices to show that for an arbitrary finite $I \subset \mathcal{I}$ the collection $\{X_i\}_{i \in I}$ is independent if, and only if, equation [\(6.5\)](#) holds.

(\Leftarrow) Let $\mathcal{C}_i := \{\{\omega : X_i(\omega) \leq x\}, x \in \mathbb{R}\} := \{X_i^{-1}((-\infty, x]), x \in \mathbb{R}\}$ and note that these are subsets of Ω . Furthermore,

1. \mathcal{C}_i is a π -system since

$$\{\omega : X_i(\omega) \leq x\} \cap \{\omega : X_i(\omega) \leq y\} = \{\omega : X_i(\omega) \leq \min\{x, y\}\}.$$

2. Recall that $\sigma(\{(-\infty, x], x \in \mathbb{R}\}) = \mathcal{B}$. By Theorem 3.1 X_i is a random variable ($X_i^{-1}(\mathcal{B}) \subset \mathcal{F}$) if, and only if,

$$\{X_i^{-1}((-\infty, x]), x \in \mathbb{R}\} = \mathcal{C}_i \subset \mathcal{F}.$$

Hence,

$$\begin{aligned} \sigma(\mathcal{C}_i) &= \sigma(\{X_i^{-1}((-\infty, x]), x \in \mathbb{R}\}) \\ &= X_i^{-1}(\sigma(\{(-\infty, x], x \in \mathbb{R}\})) \text{ by Theorem 6.1} \\ &= X_i^{-1}(\mathcal{B}) := \sigma(X_i). \end{aligned}$$

Now, equation (6.5) implies that $\{\mathcal{C}_i\}_{i \in I}$ is independent collection, therefore, by Theorem 6.2, the collection $\{\sigma(\mathcal{C}_i) = \sigma(X_i)\}_{i \in I}$ is independent. Consequently, by definition, $\{X_i\}_{i \in I}$ is an independent collection of random variables.

(\implies) This follows directly from the definition of independence. ■

Remark 6.2. 1. It follows directly from Theorem 6.4 that a finite collection of random variables $\{X_i\}_{i=1}^n$ is independent if, and only if,

$$P(\cap_{i \in I} \{\omega : X_i(\omega) \leq x_i\}) = \prod_{i \in I} P(\{\omega : X_i(\omega) \leq x_i\}), \text{ for all } I \subset \{1, \dots, n\}.$$

2. If each X_i has a density $\{X_i\}_{i=1}^n$ are independent if, and only if,

$$P(\cap_{i \in I} \{\omega : X_i(\omega) \leq x_i\}) = \prod_{i \in I} \int_{(-\infty, x_i]} f_{X_i} d\lambda.$$

6.1 Random elements

The most common cases where we deal with random elements occur when the co-domain of the element is endowed with a metric, so that the co-domain is a metric space.

Definition 6.5. Let $X : (\Omega, \mathcal{F}, P) \rightarrow (T, \mathcal{T} = \sigma(\mathcal{O}))$, where \mathcal{O} are the open sets in T . Then, X is a random element if

$$X^{-1}(B) \in \mathcal{F} \text{ for all } B \in \mathcal{T}.$$

In this definition, \mathcal{T} is the collection of Borel sets of T and we write $\mathcal{B}(T)$. The following examples include definitions.

Example 6.2. Let $X : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}^k, \mathcal{B}(\mathbb{R}^k))$ where $k \in \mathbb{N}$. Then X is a random vector if $X^{-1}(B) \in \mathcal{F}$ for all $B \in \mathcal{B}(\mathbb{R}^k)$. Now, define $d_E : \mathbb{R}^k \times \mathbb{R}^k \rightarrow [0, \infty)$ as $d_E(x, y) = \left(\sum_{i=1}^k (x_i - y_i)^2 \right)^{1/2}$. It can be easily verified that d_E is a metric on \mathbb{R}^k .

Example 6.3. Let $m : \mathbb{R}^2 \rightarrow \mathbb{R}$ be given by $m(x_1, x_2) = \frac{|x_1 - x_2|}{1 + |x_1 - x_2|}$. Clearly, from the definition of m , $m \geq 0$, $m = 0$ if, and only if, $x_1 = x_2$ and $m(x_1, x_2) = m(x_2, x_1)$. To verify that $m(x_1, x_2) \leq m(x_1, z) + m(z, x_2)$ we note that $|x_1 - x_2| = \frac{m(x_1, x_2)}{1 - m(x_1, x_2)}$. Since $|x_1 - x_2| = |x_1 - z + z - x_2| \leq |x_1 - z| + |z - x_2|$, we have

$$\frac{m(x_1, x_2)}{1 - m(x_1, x_2)} \leq \frac{m(x_1, z)}{1 - m(x_1, z)} + \frac{m(z, x_2)}{1 - m(z, x_2)}.$$

Let $c = m(x_1, x_2)$, $a = m(x_1, z)$ and $b = m(z, x_2)$. Then, $\frac{c}{1-c} \leq \frac{a}{1-a} + \frac{b}{1-b} = \frac{a+b-2ab}{(1-a)(1-b)}$ and

$$a + b \geq \frac{c}{1-c}(1-a)(1-b) + 2ab = -\frac{c}{1-c}(a+b) + \frac{c}{1-c} + \frac{1}{1-c}(2ab - abc).$$

Then,

$$\frac{a+b}{1-c} \geq \frac{c}{1-c} + \frac{1}{1-c}(2ab - abc) \iff a + b \geq c + ab(2-c).$$

Since $0 \leq m \leq 1$, $ab(2-c) \geq 0$ and $c \leq a+b$. Hence, $m(x_1, x_2) \leq m(x_1, z) + m(z, x_2)$. This shows that m is a metric on \mathbb{R}^2 .

Now, consider a space of sequences $\{x_i\}_{i \in \mathbb{N}}$ where $x_i \in \mathbb{R}$ for all i and define $m_\infty : \mathbb{R}^\infty \times \mathbb{R}^\infty \rightarrow \mathbb{R}$ as $m_\infty(\{x_i\}_{i \in \mathbb{N}}, \{y_i\}_{i \in \mathbb{N}}) = \lim_{n \rightarrow \infty} \sum_{j=1}^n \frac{1}{2^j} m(x_j, y_j) = \lim_{n \rightarrow \infty} S_n$. Since $0 \leq S_1 \leq S_2 \leq \dots$ is a monotonic sequence, it converges if, and only if, it is bounded. Boundedness follows from the fact that $|S_n| \leq \sum_{j=1}^n \frac{1}{2^j} m(x_j, y_j) \leq \sum_{j=1}^n \frac{1}{2^j} \leq \sum_{j=1}^\infty \frac{1}{2^j} = 1$. Hence, the limit in the definition of m_∞ exists and $0 \leq m_\infty \leq 1$. If $m_\infty(\{x_i\}_{i \in \mathbb{N}}, \{y_i\}_{i \in \mathbb{N}}) = 0$ then it must be that $m(x_j, y_j) = 0$ for all j , which implies $x_j = y_j$ for all j . Clearly, if $x_j = y_j$ for all j we have $m_\infty(\{x_i\}_{i \in \mathbb{N}}, \{y_i\}_{i \in \mathbb{N}}) = 0$.

Since $m(x_j, y_j) \leq m(x_j, z_j) + m(z_j, y_j)$ we have

$$\sum_{j=1}^n 2^{-j} m(x_j, y_j) \leq \sum_{j=1}^n 2^{-j} m(x_j, z_j) + \sum_{j=1}^n 2^{-j} m(z_j, y_j).$$

Taking limits on both sides as $n \rightarrow \infty$ gives $m_\infty(\{x_i\}_{i \in \mathbb{N}}, \{y_i\}_{i \in \mathbb{N}}) \leq m_\infty(\{x_i\}_{i \in \mathbb{N}}, \{z_i\}_{i \in \mathbb{N}}) + m_\infty(\{z_i\}_{i \in \mathbb{N}}, \{y_i\}_{i \in \mathbb{N}})$. Hence, m_∞ is a metric in the space of infinite sequences.

Alternatively, we can define $\mu_\infty : \mathbb{R}^\infty \times \mathbb{R}^\infty \rightarrow \mathbb{R}$ as

$$\begin{aligned} \mu_\infty(\{x_i\}_{i \in \mathbb{N}}, \{y_i\}_{i \in \mathbb{N}}) &= \lim_{n \rightarrow \infty} \sum_{j=1}^n \frac{1}{2^j} \frac{\sum_{i=1}^j |x_i - y_i|}{1 + \sum_{i=1}^j |x_i - y_i|} = \lim_{n \rightarrow \infty} \sum_{j=1}^n \frac{1}{2^j} u_j(\{x_i\}_{i=1}^j, \{y_i\}_{i=1}^j) \\ &= \lim_{n \rightarrow \infty} S_n. \end{aligned}$$

As in the case of m_∞ , $0 \leq S_1 \leq S_2 \leq \dots$ and $|S_n| \leq 1$. Hence, $0 \leq \mu_\infty \leq 1$ and $\mu_\infty = 0$ if, and only if, $x_i = y_i$ for all i . Now, write $S_1(x_1, y_1) = \frac{1}{2} m(x_1, y_1)$ and since $m(x_1, y_1) \leq m(x_1, z_1) + m(z_1, y_1)$ we have that $S_1(x_1, y_1) \leq S_1(x_1, z_1) + S_1(z_1, y_1)$. Now, suppose

$$S_n(\{x_i\}_{i=1}^n, \{y_i\}_{i=1}^n) \leq S_n(\{x_i\}_{i=1}^n, \{z_i\}_{i=1}^n) + S_n(\{z_i\}_{i=1}^n, \{y_i\}_{i=1}^n).$$

Then,

$$\begin{aligned}
S_{n+1}(\{x_i\}_{i=1}^{n+1}, \{y_i\}_{i=1}^{n+1}) &= \sum_{j=1}^{n+1} \frac{1}{2^j} u_j(\{x_i\}_{i=1}^j, \{y_i\}_{i=1}^j) = S_n(\{x_i\}_{i=1}^n, \{y_i\}_{i=1}^n) \\
&\quad + \frac{1}{2^{n+1}} u_{n+1}(\{x_i\}_{i=1}^{n+1}, \{y_i\}_{i=1}^{n+1}) \\
&\leq S_n(\{x_i\}_{i=1}^n, \{z_i\}_{i=1}^n) + S_n(\{z_i\}_{i=1}^n, \{y_i\}_{i=1}^n) \\
&\quad + \frac{1}{2^{n+1}} u_{n+1}(\{x_i\}_{i=1}^{n+1}, \{y_i\}_{i=1}^{n+1})
\end{aligned}$$

Following the same arguments used for m , we have

$$\begin{aligned}
\frac{1}{2^{n+1}} u_{n+1}(\{x_i\}_{i=1}^{n+1}, \{y_i\}_{i=1}^{n+1}) &\leq \frac{1}{2^{n+1}} u_{n+1}(\{x_i\}_{i=1}^{n+1}, \{z_i\}_{i=1}^{n+1}) + \frac{1}{2^{n+1}} u_{n+1}(\{z_i\}_{i=1}^{n+1}, \{y_i\}_{i=1}^{n+1}) \\
&= \frac{1}{2^{n+1}} \frac{\sum_{i=1}^{n+1} |x_i - z_i|}{1 + \sum_{i=1}^{n+1} |x_i - z_i|} + \frac{1}{2^{n+1}} \frac{\sum_{i=1}^{n+1} |z_i - y_i|}{1 + \sum_{i=1}^{n+1} |z_i - y_i|}.
\end{aligned}$$

Hence,

$$\begin{aligned}
S_{n+1}(\{x_i\}_{i=1}^{n+1}, \{y_i\}_{i=1}^{n+1}) &\leq S_n(\{x_i\}_{i=1}^n, \{z_i\}_{i=1}^n) + S_n(\{z_i\}_{i=1}^n, \{y_i\}_{i=1}^n) \\
&\quad + \frac{1}{2^{n+1}} \frac{\sum_{i=1}^{n+1} |x_i - z_i|}{1 + \sum_{i=1}^{n+1} |x_i - z_i|} + \frac{1}{2^{n+1}} \frac{\sum_{i=1}^{n+1} |z_i - y_i|}{1 + \sum_{i=1}^{n+1} |z_i - y_i|} \\
&= S_{n+1}(\{x_i\}_{i=1}^{n+1}, \{z_i\}_{i=1}^{n+1}) + S_{n+1}(\{z_i\}_{i=1}^{n+1}, \{y_i\}_{i=1}^{n+1}).
\end{aligned}$$

Hence, by induction, and taking limits we have $\mu_\infty(\{x_i\}_{i \in \mathbb{N}}, \{y_i\}_{i \in \mathbb{N}}) \leq \mu_\infty(\{x_i\}_{i \in \mathbb{N}}, \{z_i\}_{i \in \mathbb{N}}) + \mu_\infty(\{z_i\}_{i \in \mathbb{N}}, \{y_i\}_{i \in \mathbb{N}})$.

Example 6.4. Let $X : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}^\infty, \mathcal{B}(\mathbb{R}^\infty))$ where $\mathbb{R}^\infty = \times_{n=1}^\infty \mathbb{R}$ and $\mathcal{B}(\mathbb{R}^\infty) = \sigma(\mathcal{C})$ with $\mathcal{C} = \{C : C = \theta_i^{-1}(B), B \in \mathcal{B}^i, \theta_i(x) = (X_1, \dots, X_i) : \mathbb{R}^\infty \rightarrow \mathbb{R}^i, i \in \mathbb{N}\}$. Then X is a random sequence if $X^{-1}(B) \in \mathcal{F}$ for all $B \in \mathcal{B}(\mathbb{R}^\infty)$ and $d : \mathbb{R}^\infty \times \mathbb{R}^\infty \rightarrow [0, \infty)$ is $d(x, y) = \sum_{i=1}^\infty \frac{1}{2^i} \left(\frac{\sum_{j=1}^i |x_j - y_j|}{1 + \sum_{j=1}^i |x_j - y_j|} \right)^{1/2}$ is the metric on \mathbb{R}^∞ .

Remark 6.3. 1. Let $X \in \mathbb{R}^k$ be a random vector and $f : \mathbb{R}^k \rightarrow \mathbb{R}$ be measurable. Then, $h : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}, \mathcal{B})$ with $h(\omega) = f(X(\omega)) = (f \circ X)(\omega)$ is a random variable since compositions of measurable functions are measurable by Theorem [3.3](#). In particular the

result follows if f is continuous. That is, real valued continuous functions of random vectors are random variables.

2. In 1, if $f(X) = \pi_i(X) = X_i$ and X is random vector then X_i is a random variable for $i = 1, \dots, k$.

Theorem 6.5. $X \in \mathbb{R}^k$ is a random vector $\iff X_i$ is a random variable, where X_i is the i th component of X .

Proof. (\Leftarrow) Suppose X_i is a random variable for $i = 1, \dots, k$. Let $R_k = I_1 \times \dots \times I_k$, where $I_i = [a_i, b_i)$ are intervals in \mathbb{R} . Then,

$$\begin{aligned} X^{-1}(R_k) &= \{\omega : X_i(\omega) \in [a_i, b_i) \forall i\} \\ &= \{\omega : X_i^{-1}([a_i, b_i)) \forall i\} = \bigcap_{i=1}^k X_i^{-1}(I_i). \end{aligned}$$

Since X_i is a random variable, $X_i^{-1}(I_i) \in \mathcal{F}$. Furthermore, since \mathcal{F} is a σ -algebra, it is closed under intersections, and $X^{-1}(R_k) \in \mathcal{F}$. The other direction of the equivalence follows from the previous remark. ■

Remark 6.4. 1. Theorem [6.5](#) extends to $X = \{X_1, X_2, \dots\}$. That is, X is a random sequence if, and only if, each X_i is a random variable. Furthermore, X is a random sequence if, and only if, $(X_1 \cdots X_k)$ is random vector for any k .

2. $X^{-1}((-\infty, a_1] \times \dots \times (-\infty, a_k]) \in \mathcal{F}$ and we write $P(X^{-1}((-\infty, a_1] \times \dots \times (-\infty, a_k])) = P \circ X^{-1}(\times_{i=1}^k (-\infty, a_i]) = P_X(\times_{i=1}^k (-\infty, a_i])$.

Also, if there exists a non-negative Borel measurable function $f_X : \mathbb{R}^k \rightarrow \mathbb{R}$ that satisfies

$$P_X(\times_{i=1}^k (-\infty, a_i]) = \int_{C(a)} f_X d\lambda^k,$$

where $C(a) = \times_{i=1}^k (-\infty, a_i]$ and $a = (a_1 \cdots a_k)^T$, we call f_X the “joint density” of X . Naturally, the joint distribution function associated with X is

$$F_X(a) : \mathbb{R}^k \rightarrow [0, 1],$$

where $F_X(a) = P(C(a))$ for $a \in \mathbb{R}^k$. We can write $C(a) = \cap_{i=1}^k \{\omega : X_i(\omega) \leq a_i\}$. That $\{\omega : X_i(\omega) \leq a_i\}$ is an element of \mathcal{F} follows from Theorem [6.5](#).

Theorem 6.6. Consider two random variables $X_1, X_2 : (\Omega, \mathcal{F}, P) \rightarrow (\mathbb{R}, \mathcal{B})$. X_1 and X_2 are independent if, and only if, one of the following holds:

- a) $P(\{X_1 \in A_1\} \cap \{X_2 \in A_2\}) := P(X \in A_1, X \in A_2) = P(X_1 \in A_1)P(X_2 \in A_2)$, for all $A_1, A_2 \in \mathcal{B}$,
- b) $P(X_1 \in A_1, X_2 \in A_2) = P(X_1 \in A_1)P(X_2 \in A_2)$, for all $A_1 \in \mathcal{A}_1, A_2 \in \mathcal{A}_2$, where $\mathcal{A}_1, \mathcal{A}_2$ are π systems which generate \mathcal{B} ,
- c) $f(X_1)$ and $g(X_2)$ are independent for each pair (f, g) of measurable functions,
- d) $E(f(X_1), g(X_2)) = E(f(X_1))E(g(X_2))$ for each pair of (f, g) of bounded measurable (or non-negative measurable) functions.

Proof. First, note that X_1 and X_2 independent means that $\sigma(X_1) = X_1^{-1}(\mathcal{B})$ and $\sigma(X_2) = X_2^{-1}(\mathcal{B})$ are independent. That is, for all $A_1, A_2 \in \mathcal{B}$,

$$\begin{aligned} P(X_1^{-1}(A_1) \cap X_2^{-1}(A_2)) &= P(X_1^{-1}(A_1))P(X_2^{-1}(A_2)) \\ &\iff P(X_1 \in A_1, X_2 \in A_2) = P(X_1 \in A_1)P(X_2 \in A_2). \end{aligned}$$

[a) \implies b)] Since \mathcal{A}_1 generates \mathcal{B} and \mathcal{A}_2 generates \mathcal{B} , $\mathcal{A}_1 \subset \mathcal{B}$ and $\mathcal{A}_2 \subset \mathcal{B}$, and if a) is true for all $A_1 \in \mathcal{B}, A_2 \in \mathcal{B}$, then b) is true.

[b) \implies a)] Let $C_1 = \{A \in \mathcal{B} : P(X_1 \in A, X_2 \in A_2) = P(X_1 \in A)P(X_2 \in A_2)\}$ for a given $A_2 \in \mathcal{A}_2$. From the proof of Theorem [6.2](#), C_1 is a Dynkin system. $\mathcal{A}_1 \subset C_1$ and $\delta(\mathcal{A}_1) = \sigma(\mathcal{A}_1) = \mathcal{B} \subset C_1$. Analogously, $C_2 = \{A \in \mathcal{B} : P(X_1 \in A_1, X_2 \in A) = P(X_1 \in A_1)P(X_2 \in A)\}$ for a given $A_1 \in \mathcal{A}_1$ is such that $\delta(\mathcal{A}_2) = \sigma(\mathcal{A}_2) = \mathcal{B} \subset C_2$. Consequently, b) \implies a).

[c) \implies a)] The identity function is measurable, therefore take $f(x) = g(x) = x$

[a) \implies c)] For concreteness, let $f : (\mathbb{R}, \mathcal{B}) \rightarrow (M_f, \mathcal{M}_f)$ and $g : (\mathbb{R}, \mathcal{B}) \rightarrow (M_g, \mathcal{M}_g)$. f measurable implies that for all $M \in \mathcal{M}_f$, $f^{-1}(M) \in \mathcal{B}$. But X_1 a random variable implies that $X_1^{-1}(f^{-1}(M)) \in \mathcal{F}$ which we can write as $(X_1^{-1} \circ f^{-1})(M) \in \mathcal{F}$. In addition, $X_1^{-1}(f^{-1}(M)) := (X_1^{-1} \circ f^{-1})(M) \in X_1^{-1}(\mathcal{B})$. Analogously, $X_2^{-1}(g^{-1}(M')) = X_2^{-1} \circ g^{-1}(M') \in X_2^{-1}(\mathcal{B})$, for all $M' \in \mathcal{M}_g$. But by a) $X_1^{-1}(\mathcal{B})$ and $X_2^{-1}(\mathcal{B})$ are independent. Therefore $f(X_1)$ and $g(X_2)$ are independent.

[d) \implies a)] Let $f = I_{A_1}$ and $g = I_{A_2}$. Then,

$$f(X_1) = \begin{cases} 1 & \text{if } X_1 \in A_1 \\ 0 & \text{if } X_1 \notin A_1 \end{cases} \text{ and } g(X_2) = \begin{cases} 1 & \text{if } X_2 \in A_2 \\ 0 & \text{if } X_2 \notin A_2. \end{cases}$$

with $E(f(X_1)) = P(X_1 \in A_1)$ and $E(g(X_2)) = P(X_2 \in A_2)$. By d)

$$E(f(X_1)g(X_2)) = P(\{X_1 \in A_1\} \cap \{X_2 \in A_2\}) = P(X_1 \in A_1)P(X_2 \in A_2).$$

Hence, d) \implies a).

[a) \implies d)] From the implication [d) \implies a)] we see that if f, g are indicator functions in d) $E(f(X_1)g(X_2)) = P(\{X_1 \in A_1\} \cap \{X_2 \in A_2\})$, which by independence a) is $P(X_1 \in A_1)P(X_2 \in A_2) = E(f(X_1))E(g(X_2))$.

Now, suppose f and g are simple functions of X_1 and X_2 . Then,

$$\begin{aligned} f(X_1) &= \sum_{i=0}^{k_f} a_i^f I_{\{X_1 \in A_i^f\}} \text{ and } E(f(X_1)) = \sum_{i=0}^{k_f} a_i^f P(X_1 \in A_i^f), \\ g(X_2) &= \sum_{i=0}^{k_g} a_i^g I_{\{X_2 \in A_i^g\}} \text{ and } E(g(X_2)) = \sum_{i=0}^{k_g} a_i^g P(X_2 \in A_i^g) \end{aligned}$$

Consequently,

$$\begin{aligned} E(f(X_1)g(X_2)) &= E \left(\sum_{i=0}^{k_f} \sum_{j=0}^{k_g} a_i^f a_j^g I_{\{X_1 \in A_i^f\} \cap \{X_2 \in A_j^g\}} \right) \\ &= \sum_{i=0}^{k_f} \sum_{j=0}^{k_g} a_i^f a_j^g P(X_1 \in A_i^f) P(X_2 \in A_j^g) \text{ by independence} \\ &= E(f(X_1))E(g(X_2)) \end{aligned} \tag{6.6}$$

Now, let f be a measurable non-negative function such that $\{f_n\}_{n \in \mathbb{N}}$ are simple functions increasing to f and g is non-negative and simple. Then,

$$\begin{aligned}
 E(f(X_1)g(X_2)) &= E\left(\lim_{n \rightarrow \infty} f_n(X_1)g(X_2)\right) \\
 &= \lim_{n \rightarrow \infty} E(f_n(X_1)g(X_2)) \text{ by Lebesgue's Monotone Convergence Theorem} \\
 &= \lim_{n \rightarrow \infty} E(f_n(X_1))E(g(X_2)) \text{ by equation (6.6)} \\
 &= E(f(X_1))E(g(X_2)) \text{ by Lebesgue's Monotone Convergence Theorem}
 \end{aligned} \tag{6.7}$$

Now, let f be non-negative and let $\{g_n\}_{n \in \mathbb{N}}$ be non-negative simple functions increasing to g measurable and non-negative. Then,

$$\begin{aligned}
 E(f(X_1)g(X_2)) &= E\left(f(X_1) \lim_{n \rightarrow \infty} g_n(X_2)\right) \\
 &= \lim_{n \rightarrow \infty} E(f(X_1)g_n(X_2)) \\
 &= \lim_{n \rightarrow \infty} E(f(X_1))E(g_n(X_2)) \text{ by equation (6.7)} \\
 &= E(f(X_1))E(g(X_2))
 \end{aligned}$$

Finally, let $f = f^+ - f^-$ be bounded and measurable and g bounded and non-negative.

$$\begin{aligned}
 E(f(X_1)g(X_2)) &= E([f^+(X_1) - f^-(X_1)]g(X_2)) \\
 &= E(f^+(X_1)g(X_2)) - E(f^-(X_1)g(X_2)) \\
 &= E(f^+(X_1))E(g(X_2)) - E(f^-(X_1))E(g(X_2)) \\
 &= E(f(X_1))E(g(X_2)).
 \end{aligned}$$

To complete the proof, repeat the last argument for $g = g^+ - g^-$. ■

6.2 Exercises

1. Let $X_1, X_2 \in \mathcal{L}^2$ and define $\text{Cov}(X_1, X_2) = E([X_1 - E(X_1)][X_2 - E(X_2)])$. Show that $\text{Cov}(X_1, X_2) = E(X_1X_2) - E(X_1)E(X_2)$ and that if X_1 and X_2 are independent $\text{Cov}(X_1, X_2) = 0$.

2. Let $\{X_n\}_{n \in \mathbb{N}}$ be a sequence of random variables that are independent and share the same continuous distribution. Let p be a permutation of $\{1, \dots, n\}$ for $n \in \mathbb{N}$. Show that (X_1, \dots, X_n) and $(X_{p(1)}, \dots, X_{p(n)})$ have the same distribution.
3. Let I be a finite index set and consider the collection of σ -algebras $\{\mathcal{B}_i\}_{i \in I}$. Show that this collection is independent if, and only if, for every choice of non-negative \mathcal{B}_i -measurable random variable X_i , we have $E\left(\prod_{i \in I} X_i\right) = \prod_{i \in I} E(X_i)$.
4. If E is an event that is independent of the π -system P and $E \in \sigma(P)$, then $P(E)$ is either 0 or 1.
5. Let $\{A_i\}_{i=1}^n$ be independent events. Show that $P\left(\cup_{i=1}^n A_i\right) = 1 - \prod_{i=1}^n P(A_i^c)$.
6. We have proved that if X and Y are independent, then $f(X)$ and $g(Y)$ are independent if f and g are measurable. Is it possible to have X and Y are dependent and $f(X)$ and $g(Y)$ are independent? If so, give an example, if not, prove.