

HOMEWORK 4.

1. Since $E(a_t | a_{t-1}) = 0$, $E(a_t) = 0$ and $V(a_t | a_{t-1}) = 1 + 0.5 a_{t-1}^2$
 $\underline{=}$
 $V(a_t) = 1 + 0.5 E(a_{t-1}^2)$.

• If $\{a_t\}$ is covariance stationary, then $V(a_t) = \frac{1}{1 - 0.5}$

• From the class notes, if $E(a_t | a_{t-1}) = 0$, then $\text{cov}(a_t, a_{t-h}) =$

$$\gamma(h) = 0 \quad \text{for } h = 1, 2, \dots$$

Hence, r_t is an AR(1) process with noise $\{a_t\}$ such that

$E(a_t) = 0$, $V(a_t) = 2$ and $\gamma(h) = 0$ for $h > 1$. Hence

$$(a) E(r_t) = \frac{3}{1 - 0.7} = 10 \text{ since } |0.7| < 1.$$

$$(b) V(r_t) = \frac{V(a_t)}{1 - (0.7)^2} = 3.92$$

$$(c) E(r_t - E(r_t))(r_{t-1} - E(r_{t-1})) = 0.7 V(r_t)$$

$$E(r_t - E(r_t))(r_{t-k} - E(r_{t-k})) = (0.7)^k V(r_t)$$

$$d) \quad a_t = (\alpha_0 + \alpha_1 a_{t-1}^2)^{\frac{1}{2}} \epsilon_t$$

$$E(a_t^2 | a_{t-1}, \dots) = \alpha_0 + \alpha_1 a_{t-1}^2$$

Now,

$$\begin{aligned} a_t^2 &= \alpha_0 + \alpha_1 a_{t-1}^2 + a_t^2 - E(a_t^2 | a_{t-1}, \dots) \\ &= \alpha_0 + \alpha_1 a_{t-1}^2 + v_t \end{aligned}$$

$$\text{where } v_t = a_t^2 - E(a_t^2 | a_{t-1}, \dots) = (\alpha_0 + \alpha_1 a_{t-1}^2)(\epsilon_t^2 - 1)$$

Note that $\{a_t^2\}$ has an AR(1) structure with noise v_t .

$$E(v_t | a_{t-1}, \dots) = (\alpha_0 + \alpha_1 a_{t-1}^2) E(\epsilon_t^2 - 1) = 0. \text{ Consequently}$$

$$E(v_t) = E(E(v_t | a_{t-1}, \dots)) = 0 \text{ and } E(v_t(h)) = 0 \text{ for } h = 1, 2, \dots$$

since the conditional expectation $E(v_t | a_{t-1}, \dots) = 0$.

$$\begin{aligned} E(v_t^2) &= V(v_t) = E(E(v_t^2 | a_{t-1})) \\ &= E((\alpha_0 + \alpha_1 a_{t-1}^2)^2 E((\epsilon_t^2 - 1)^2)) \\ &= E((\alpha_0 + \alpha_1 a_{t-1}^2)^2) E((\epsilon_t^2 - 1)^2) \end{aligned}$$

$$\begin{aligned} E((\alpha_0 + \alpha_1 a_{t-1}^2)^2) &= E(\alpha_0^2 + 2\alpha_0 \alpha_1 a_{t-1}^2 + \alpha_1^2 a_{t-1}^4) \\ &= \alpha_0^2 + 2\alpha_0 \alpha_1 E(a_{t-1}^2) + \alpha_1^2 E(a_{t-1}^4) \\ &= \alpha_0^2 + 2\alpha_0 \alpha_1 \frac{\alpha_0}{1 - \alpha_1} + \alpha_1^2 E(a_{t-1}^4) \\ &= \alpha_0^2 + \frac{2\alpha_0 \alpha_1^2}{1 - \alpha_1} + \alpha_1^2 E(a_{t-1}^4) \end{aligned}$$

$$\text{Thus, } V(v_t) = \left(\alpha_0^2 + \frac{2\alpha_0 \alpha_1^2}{1 - \alpha_1} + \alpha_1^2 E(a_{t-1}^4) \right) V(\epsilon_t^2)$$

If $V(\epsilon_t^2)$, $E(a_{t-1}^4)$ exist $V(v_t) < \infty$ hence, a_t^2 is an AR(1) with noise v_t which is uncorrelated with

variance of $V(v_t)$. Consequently, from our work with

AR(1) processes

$$r_{at}^2(h) = \frac{1}{(1 - 0.5^2)} 0.5^h V(v_t)$$

Q. $r_1 - 0.7 = \phi(r_0 - 0.7) + a_1$
 $0.3 = -0.25 + a_1 \Rightarrow a_1 = 0.55$

$$\begin{aligned} r_2 - 0.7 &= 0.5(1 - 0.7) + a_2 \\ &= 0.15 + \varepsilon_2 (1 + 0.3 a_1^2)^{1/2} \\ \Rightarrow r_2 &= 0.85 + 1.0908 \varepsilon_2 \end{aligned}$$

$$\begin{aligned} E(r_2 | r_0 = 0.2, r_1 = 1) &= 0.85 \\ V(r_2 | r_0 = 0.2, r_1 = 1) &= 1.0908 \times V(\varepsilon_2) = 1.0908 \end{aligned}$$

3.a) For (a) use code `arma-garch-apple.m`, model 1 on
code lines 21 to 24.

(b) use the same code as (a), but model 2 on lines
27 to 30. On line 28, change "distS" to "distg"

(c) Use model 2 on the same code, but change ('ARLags', 1)
to ('ARLags', 2) and ('GARCHLags', 1) to ('GARCHLags', 2)

b) Just verify if the Tstatistic associated with GARCH(1,1)³ on the MATLAB output is greater than or equal to 1.96. If so, reject $H_0: \beta_1 = 0$.

c) Same as b), but look at the t-statistic at the ARIMA table from MATLAB. Compare to the 90% quantile of a normally distributed random variable.

4// a) Use code garch-student1.m, model 1, lines 28-36
 b) " " " , model 2, lines 38-46

5// a), b), c) answers are directly from class notes
 (as derived on the board) with $N=2$. That is,

a) $\hat{w} = \bar{w}/c_p$, c_p and \bar{w} are defined in your class notes

notes

$$b) \hat{w}_f = 1 - \sum_{i=1}^2 \hat{w}_i$$

$$c) w_T = \bar{w} / \sum_{i=1}^2 \bar{w}_i$$

$$d) \hat{\mu}_1 = \frac{1}{n} \sum_{t=1}^n R_{t,A}, \quad \hat{\mu}_2 = \frac{1}{n} \sum_{t=1}^n R_{t,B}$$

$$\text{Cov}(R_{t,A}, R_{t,B}) = \frac{1}{n} \sum_{t=1}^n (R_{tA} - \hat{\mu}_1) (R_{tB} - \hat{\mu}_2)$$

$$V(R_{t,A}) = \frac{1}{n} \sum_{t=1}^n (R_{tA} - \hat{\mu}_1)^2$$

$$V(R_{t,B}) = \frac{1}{n} \sum_{t=1}^n (R_{tB} - \hat{\mu}_2)^2$$

e) Yes. Since $\{R_{t,A}\}$ is assumed to be IID.

$$E(\hat{\mu}_1) = \frac{1}{n} n \cdot \mu_1 = \mu_1$$